

Meta-model for sum of indicator random variables

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Outline

- 1 Introduction
- 2 Chaos decomposition of the indicator function
- 3 PCE for a sum of indicator functions
- 4 Meta-model for a Large sum of indicator functions

Introduction

Subject of study

- ▶ We study \mathcal{L}_K , a random variable taking values in \mathbb{R} such that

$$\mathcal{L}_K = \sum_{k=1}^K \mathbf{1}_{c_k \leq X}$$

with $K \in \mathbb{N}^*$, X random variable with law ν and $(c_k)_{k \in \mathbb{N}^*}$ a sequence of independent random variables and independent from X .

Motivations

- ▶ Simulation methods (Monte Carlo) for the computation of : $\mathbb{E}[f(\mathcal{L}_K)]$, quantiles, etc...
- ▶ If K is large, the simulation cost of \mathcal{L}_K is high.

Introduction

Objectives

- ▶ Propose a polynomial chaos expansion (truncated to order $N \in \mathbb{N}$) for the indicator and so for \mathcal{L}_K :

$$\mathbf{1}_{c_k \leq X} = \sum_{n=0}^{+\infty} \gamma_n(c_k) p_n(X) \approx \sum_{n=0}^N \gamma_n(c_k) p_n(X), \quad \mathcal{L}_{K,N} = \sum_{k=1}^K \sum_{n=0}^N \gamma_n(c_k) p_n(X)$$

where $(p_n)_{n \in \mathbb{N}}$ is an Orthogonal Polynomial Set (OPS) for ν :

$$\mathbb{E}[p_n(X)p_m(X)] = \int_{\mathbb{R}} p_n(x)p_m(x)\nu(dx) = h_n \mathbf{1}_{n=m}.$$

Interest of the chaos decompositions

- ▶ Gaussian Approximation $\gamma_{K,n}^G$ of $\sum_{k=1}^K \gamma_n(c_k)$ which satisfies a **CLT** when c_k are independent r.v. and $K \rightarrow \infty$:

$$\mathcal{L}_{K,N}^G = \sum_{n=0}^N \gamma_{K,n}^G p_n(X)$$

with $(\gamma_{K,n}^G)_{n=0,\dots,N}$ a **Gaussian** vector and $N \ll K$. Parameters computed only **one time** and not for each simulation.

Introduction

Organisation of the talk

- ▶ Chaos decomposition of the indicator function $\mathbf{1}_{c \leq X}$.
- ▶ Chaos decomposition (truncated) $\mathcal{L}_{K,N} = \sum_{k=1}^K \sum_{n=0}^N \gamma_n(c_k) p_n(X)$ of $\mathcal{L}_K = \sum_{k=1}^K \mathbf{1}_{c_k \leq X}$.
- ▶ Meta-model $\mathcal{L}_{K,N}^{\mathcal{G}} = \sum_{n=0}^N \gamma_n^{\mathcal{G}} p_n(X)$ for \mathcal{L}_K in the Gaussian Copula model

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Indicator function

Chaos decomposition

- ▶ Consider a probability measure ν .
- ▶ Condition of existence of the OPS $(p_n)_{n \in \mathbb{N}}$ for ν : There exists $\kappa > 0$ such that

$$\int_{\mathbb{R}} e^{\kappa|x|} \nu(dx) < +\infty,$$

Under this condition

- $(p_n)_{n \in \mathbb{N}}$ is orthogonal for ν
 - $h_n := \|p_n\|_{L^2(\mathbb{R}, \nu)}^2 = \int_{\mathbb{R}} p_n(x) p_n(x) \nu(dx) < +\infty,$
 - $(p_n)_{n \in \mathbb{N}}$ is dense in $L^2(\mathbb{R}, \nu)$
- ▶ For every $f \in L^2(\nu)$, the following Polynomial Chaos Expansion holds

$$f \stackrel{L^2(\nu)}{=} \sum_{n=0}^{+\infty} \gamma_n^f p_n, \quad \gamma_n^f = \frac{\langle f, p_n \rangle_{L^2(\nu)}}{\|p_n\|_{L^2(\nu)}^2}.$$

Indicator function

Preliminaries

- Framework : c deterministic in \mathbb{R} and X random variable in \mathbb{R} with law ν such that $\nu(dx) = w(x)\mathbf{1}_{I_w}(x)dx$. **Classical Orthogonal Polynomial Set :**

COPS	Hermite	Laguerre	Jacobi	Legendre
Param.		$\alpha > -1$	$\alpha, \beta > -1$	$\alpha = \beta = 0$
ν	$\mathcal{N}(0, 1)$	$\text{Gamma}(\alpha + 1, 1)$	$1 - 2\text{Beta}(\alpha + 1, \beta + 1)$	$\mathcal{U}(-1, 1)$
$p_n(x)$	$\text{He}_n(x)$	$L_n(x)$	$J_n(x)$	$\text{Le}_n(x)$
I_w	$(-\infty, +\infty)$	$(0, +\infty)$	$(-1, 1)$	$(-1, 1)$
$w(x)$	$\frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}}$	$\frac{x^\alpha e^{-x}}{\Gamma_{\alpha+1}}$	$\frac{(1-x)^\alpha (1+x)^\beta}{2^{\alpha+\beta+1} \text{B}_{\alpha+1, \beta+1}}$	$\frac{1}{2}$

- Objective : For $c \in I_w$, study the PCE

$$\mathbf{1}_{c \leq X} = \sum_{n=0}^{\infty} \gamma_n(c) p_n(X).$$

Indicator function

Choice of polynomial basis

Let X with law ν and denote $\Psi_\nu(c) = \mathbb{P}(X \leq c)$. Let ν_1 be a COPS measure, we denote $T_{\nu, \nu_1} = \Psi_{\nu_1}^{-1} \circ \Psi_\nu$, and then $X_{\nu_1} := T_{\nu, \nu_1}(X) \sim \nu_1$

$$\mathbf{1}_{c \leq X} = \mathbf{1}_{T_{\nu, \nu_1}(c) \leq X_{\nu_1}}$$

Denoting $q = \Psi_\nu(c)$,

$$\mathbf{1}_{\Psi_\nu^{-1}(q) \leq X} = \mathbf{1}_{\Psi_{\nu_1}^{-1}(q) \leq X_{\nu_1}}$$

Indicator function

Proposition - PCE for the indicator function of a COPS

Let $(p_n)_{n \in \mathbb{N}}$ be a COPS w.r.t. $\nu(dx) = w(x)\mathbf{1}_{I_w}(x)dx$, and let $c \in I_w$. Then,

$$\forall x \in I_w \setminus \{c\}, \quad \mathbf{1}_{c \leq x} = \sum_{n=0}^{+\infty} \gamma_n(c) p_n(x),$$

with $\gamma_0(c) = \nu((c, +\infty))$ for every $n \geq 1$,

$$\gamma_n(c) = \frac{e^{-\frac{c^2}{2}} \text{He}_{n-1}(c)}{n! \sqrt{2\pi}} \quad \text{if } \nu = \mathcal{N}(0, 1)$$

$$\gamma_n(c) = -\frac{(n-1)!}{\Gamma_{n+\alpha+1}} c^{\alpha+1} e^{-c} L_{n-1}^{(\alpha+1)}(c) \quad \text{if } \nu = \text{Gamma}(\alpha+1, 1)$$

$$\gamma_n(c) = C_n(\alpha, \beta) (1-c)^{\alpha+1} (1+c)^{\beta+1} J_{n-1}^{(\alpha+1, \beta+1)}(c) \quad \text{if } \nu = 1 - 2\text{Beta}(\alpha+1, \beta+1)$$

$$\text{with } C_n(\alpha, \beta) = \frac{(2n+\alpha+\beta+1)\Gamma_{n+\alpha+\beta+1}(n-1)!}{2^{\alpha+\beta+2}\Gamma_{n+\alpha+1}\Gamma_{n+\beta+1}}.$$

Idea of the proof : Use the Rodrigue's formula $p_n(x) = \frac{1}{\kappa_n w(x)} \frac{d^n}{dx^n} [F(x)^n w(x)]$ where F (pol. of degree ≤ 2) and κ_n are known explicitly for COPS.

Indicator function

For OPS, $p_{n+1}(x) = (A_n x + B_n) p_n(x) - C_n p_{n-1}(x)$, $n \in \mathbb{N}^*$ where A_n, B_n, C_n are constants with $A_n > 0, C_n > 0$.

Proposition - Recurrence relation for the Chaos coefficient

Let $(p_n)_{n \in \mathbb{N}}$ be a COPS w.r.t. $\nu(dx) = w(x)1_{I_w}(x)dx$ and $c \in I_w = (a, b)$, then

$$\begin{aligned} \gamma_{n+2}(c) &= (B_{n+1} + A_{n+1}c) \frac{h_{n+1}}{h_{n+2}} \gamma_{n+1}(c) - C_{n+1} \frac{h_n}{h_{n+2}} \gamma_n(c) \\ &\quad + \frac{h_{n+1}}{h_{n+2}} A_{n+1} \int_c^b \gamma_{n+1}(x) dx. \end{aligned}$$

Moreover, there exist $D_{\nu,n}, E_{\nu,n}, F_{\nu,n}$ ($D_{\nu,n}, F_{\nu,n} \geq 0$) such that

$$\gamma_{n+2}(c) = (D_{\nu,n} + cE_{\nu,n})\gamma_{n+1}(c) - F_{\nu,n}\gamma_n(c),$$

Indicator function - L_2 truncation error

Let us define

$$\mathcal{E}_N(c) := \mathbb{E} \left[\left| \mathbf{1}_{c \leq X} - \sum_{n=0}^N \gamma_n(c) p_n(X) \right|^2 \right]^{\frac{1}{2}} = \left| \sum_{n=N+1}^{+\infty} h_n \gamma_n(c)^2 \right|^{\frac{1}{2}}.$$

If ν is either $\mathcal{N}(0, 1)$ (Hermite), $\text{Gamma}(\alpha + 1, 1)$ (Laguerre) or $1 - 2\text{Beta}(\alpha + 1, \beta + 1)$ (Jacobi)

$$E_{\nu, N}(c) := e^{-\frac{c^2}{4}} \begin{cases} N^{-\frac{1}{4}}, & |c| \in [0, (4N)^{\frac{1}{2}}], \\ \frac{N^{1/24}}{|c|^{\frac{1}{4}}}, & |c| \in [(2N)^{\frac{1}{2}}, N^{\frac{39}{14}}], \\ \frac{|c|^{1/3}}{N^{19/12}}, & |c| \in [N^{\frac{39}{14}}, +\infty). \end{cases}$$

$$E_{\nu, N}(c) := e^{-\frac{c}{2}} \begin{cases} \frac{c^{\frac{\alpha}{2} + \frac{1}{4}}}{N^{1/4}}, & c \in (0, (4 - \eta)N], \\ \frac{c^{\frac{\alpha}{2} + \frac{5}{12}}}{N^{1/4}}, & c \in [2N, +\infty). \end{cases}$$

$$E_{\nu, N}(c) := \left(\frac{2 + \sqrt{(\alpha + 1)^2 + (\beta + 1)^2}}{2^{\alpha + \beta} \mathbf{B}_{\alpha + 1, \beta + 1}} \right)^{\frac{1}{2}} (1 - c)^{\frac{\alpha}{2} + \frac{1}{4}} (1 + c)^{\frac{\beta}{2} + \frac{1}{4}} N^{-\frac{1}{2}}, \quad c \in [-1, 1].$$

Indicator function - L_2 truncation error

Proposition - L^2 approximation

Let $(p_n)_{n \in \mathbb{N}}$ be a COPS w.r.t. $\nu(dx) = w(x)\mathbf{1}_{I_w}(x)dx$, and $c \in I_w$, then there exists $A \geq 0$ such that

$$\mathcal{E}_N(c) = \mathbb{E} \left[\left| \mathbf{1}_{c \leq X} - \sum_{n=0}^N \gamma_n(c) p_n(X) \right|^2 \right]^{\frac{1}{2}} \leq A E_{\nu, N}(c).$$

Comparison of polynomial basis

For $q = \Psi_\nu(c)$, and ν_1 a COPS measure, and $X_{\nu_1} = \Psi_{\nu_1}^{-1} \circ \Psi_\nu(X)$,

$$\mathbf{1}_{\Psi_\nu^{-1}(q) \leq X} = \mathbf{1}_{\Psi_{\nu_1}^{-1}(q) \leq X_{\nu_1}}$$

To compare two COPS, we compare their L_2 error, $\mathcal{E}_{N, \nu}(\Psi_\nu^{-1}(q))$ and $\mathcal{E}_{N, \nu_1}(\Psi_{\nu_1}^{-1}(q))$ for $q \in (0, 1)$.

Indicator function - L_2 truncation error

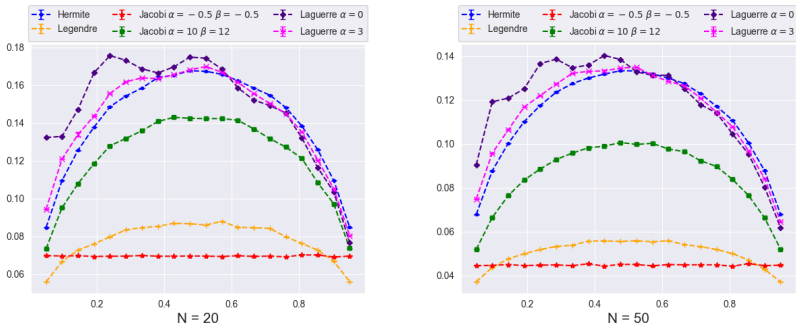


FIGURE – L^2 error $\mathcal{E}_N(c)$ for $c = \Psi_\nu^{-1}(q)$ for the COPS w.r.t. $q \in (0, 1)$ and for $N \in \{20, 50\}$.

Indicator function - L_2 truncation error

Proposition - Extreme quantile

Let $N \in \mathbb{N}$ fixed. Take $c(q) = \Psi_\nu^{-1}(q)$. We have, for the Hermite polynomials

$$E_{\nu,N}(c(q)) \sim \begin{cases} C |\ln q|^{\frac{5}{12}} q^{\frac{1}{2}}, & (q \rightarrow 0^+), \\ C |\ln(1-q)|^{\frac{5}{12}} (1-q)^{\frac{1}{2}}, & (q \rightarrow 1^-). \end{cases}$$

For the Laguerre polynomials,

$$E_{\nu,N}(c(q)) \sim \begin{cases} C q^{\frac{2\alpha+1}{4(\alpha+1)}}, & (q \rightarrow 0^+), \\ C |\ln(1-q)|^{\frac{5}{12}} (1-q)^{\frac{1}{2}}, & (q \rightarrow 1^-). \end{cases}$$

For the Jacobi polynomials,

$$E_{\nu,N}(c(q)) \sim \begin{cases} C q^{\frac{2\beta+1}{4(\alpha+1)}}, & (q \rightarrow 0^+), \\ C (1-q)^{\frac{2\alpha+1}{4(\beta+1)}}, & (q \rightarrow 1^-). \end{cases}$$

Indicator function - Extreme quantile

Best polynomials when $q \rightarrow 0^+$

- Jacobi when $\frac{2\beta_J+1}{4(\alpha_J+1)} > \frac{1}{2} \Leftrightarrow \alpha_J < \beta_J - \frac{1}{2}$,
- Hermite polynomials when $\alpha_J > \beta_J - \frac{1}{2}$,
- both Hermite and Jacobi when $\alpha_J = \beta_J - \frac{1}{2}$.

Worst L^2 estimates for Jacobi when $\frac{2\beta_J+1}{\alpha_J+1} < \frac{2\alpha_{L_a}+1}{\alpha_{L_a}+1}$.

Best polynomials when $q \rightarrow 1^-$

- Jacobi when $\frac{2\alpha_J+1}{4(\beta_J+1)} > \frac{1}{2} \Leftrightarrow \beta_J < \alpha_J - \frac{1}{2}$,
- Hermite and Laguerre when $\beta_J > \alpha_J - \frac{1}{2}$,
- All polynomials when $\beta_J = \alpha_J - \frac{1}{2}$ (up to some logarithmic error terms).

Worst L^2 estimates for Jacobi when $\alpha_J < \beta_J + \frac{1}{2}$.

Indicator function - L_2 truncation error

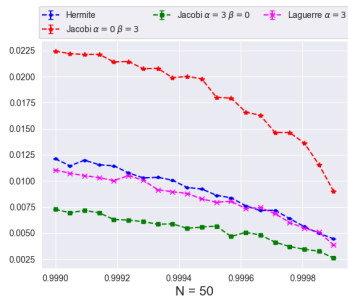
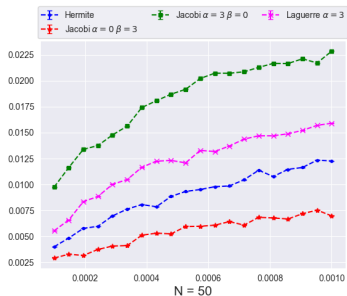


FIGURE – L^2 error for the COPS w.r.t. small quantiles $q \in (10^{-4}, 10^{-3})$ (left) and large quantiles $q \in (1 - 10^{-3}, 1 - 10^{-4})$ (right) for a fixed $N = 50$.

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Sum of Indicator functions - L_2 truncation error

We consider (introduction of deterministic weights l_k)

$$\mathcal{L}_K = \sum_{k=1}^K l_k \mathbf{1}_{c_k \leq X}, \quad \mathcal{L}_{K,N} = \sum_{k=1}^K l_k \sum_{n=0}^N \gamma_n(c_k) p_n(X)$$

$$\text{and } \mathcal{E}_N(c) = \mathbb{E} \left[\left| \mathbf{1}_{c \leq X} - \sum_{n=0}^N \gamma_n(c) p_n(X) \right|^2 \right]^{\frac{1}{2}}.$$

Proposition - L_2 truncation error

Assume $(c_k)_{k=1, \dots, K}$ is a sequence of independent random variables taking values in I_w . Then

$$\mathbb{E} \left[(\mathcal{L}_K - \mathcal{L}_{K,N})^2 \right]^{\frac{1}{2}} \leq \sum_{k=1}^K l_k^2 \mathbb{E} [\mathcal{E}_N(c_k)],$$

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Meta-model - CLT

Recall $\mathcal{L}_{K,N} = \sum_{n=0}^N \sum_{k=1}^K l_k \gamma_n(c_k) p_n(X)$.

Consider the random vector $\epsilon_K := (\epsilon_{K,n})_{n=0,\dots,N}$ for N fixed, such that

$$\epsilon_{K,n} = \sum_{k=1}^K l_k \gamma_n(c_k), \quad \text{so} \quad \mathcal{L}_{K,N} = \sum_{n=0}^N \epsilon_{K,n} p_n(X).$$

and

$$\begin{aligned} \mu_{K,n} &= \mathbb{E}[\epsilon_{K,n}] = \sum_{k=1}^K l_k \mathbb{E}[\gamma_n(c_k)], \\ \Sigma_{K,n,m} &= \text{Cov}(\epsilon_{K,n}, \epsilon_{K,m}) = \sum_{k=1}^K l_k^2 \text{Cov}(\gamma_n(c_k), \gamma_m(c_k)). \end{aligned}$$

Meta-model - CLT

Proposition - CLT

Assume that

$$\|\Sigma_K^{-1}\| \sup_{1 \leq k \leq K} l_k^2 \xrightarrow{K \rightarrow +\infty} 0 \quad \text{with} \quad \|\Sigma_K^{-1}\| := \sup_{x \in \mathbb{R}^{N+1}, x \neq 0} \frac{|\Sigma_K^{-1}x|}{|x|}.$$

Then, the vector ϵ_K satisfies the following Central Limit Theorem

$$\Sigma_K^{-1/2}(\epsilon_K - \mu_K) \xrightarrow{K \rightarrow \infty} \mathcal{N}(0, \text{Id}_{N+1}).$$

We introduce the Gaussian approximation for ϵ_K :

$$\epsilon_K^G = \left(\epsilon_{K,n}^G : n = 0, \dots, N \right) \sim \mathcal{N}(\mu_K, \Sigma_K),$$

leading to the meta-model

$$\mathcal{L}_{K,N}^G = \sum_{n=0}^N \epsilon_{K,n}^G p_n(X).$$

Application to a Gaussian Copula model

Consider

- $K = 5 \times 10^5$.
- Let $\rho = 0.1$, $\rho = 0.01$, and

$$c_k = \frac{\sqrt{1 - \rho^2}}{\rho} \varepsilon_k - \frac{1}{\rho} \Psi_{\mathcal{N}(0,1)}^{-1}(\rho).$$

- $X, \varepsilon_1, \dots, \varepsilon_K$, i.i.d. $\mathcal{N}(0, 1)$.
- $l_k = 1/\sqrt{k}$.

The Lindeberg condition holds as $\sup_{1 \leq k \leq K} l_k^2 = 1$ and $\sum_{k=1}^{+\infty} \frac{1}{k} = +\infty$.

Simulation cost : About 17 times faster (when $N = 3$) than a naive Monte Carlo.

Application to a Gaussian Copula model

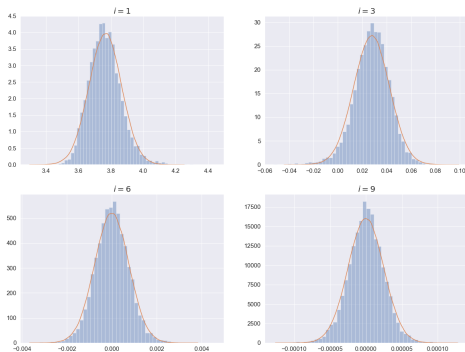


FIGURE – Histograms of $\epsilon_{K,i}$ with the p.d.f of $\mathcal{N}(\mathbb{E}[\epsilon_{K,i}], \text{Var}(\epsilon_{K,i}))$ for $N \in \{1, 3, 6, 9\}$

Application to a Gaussian Copula model

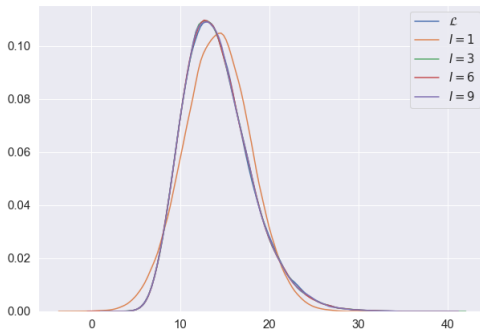


FIGURE – Distributions of $\mathcal{L}_{K,N}^G$ for $N = l \in \{1, 3, 6, 9\}$ and $\mathcal{L}_K = \mathcal{L}$

Application to a Gaussian Copula model

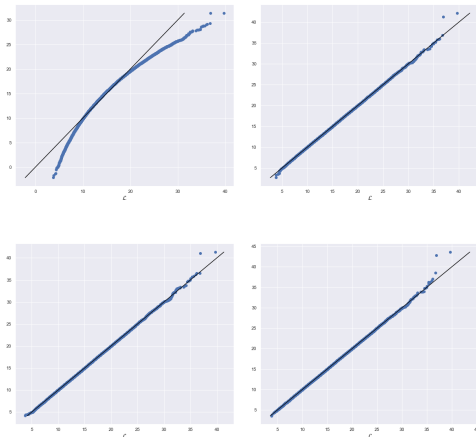


FIGURE – Q-Q plots of $\mathcal{L}_{K,N}^G$ w.r.t. \mathcal{L}_K for $N = 1, 3, 6, 9$ and 10^5 points.

Thank you for your attention !

- With Florian Bourgey and Emmanuel Gobet. *Meta-model of a large credit risk portfolio in the Gaussian copula model*, SIAM Journal on Financial Mathematics, 2020.
- With Florian Bourgey and Emmanuel Gobet. *A comparison study of polynomial-type chaos expansions for indicator functions*, 2021, Hal.