

# Analysis and optimization of certain parallel Monte Carlo methods in the low temperature limit

---

Guo-Jhen Wu\* and Paul Dupuis \*\*

RESIM 2021, May 19, 2021

\* Department of Mathematics, KTH Royal Institute of Technology

\*\* Division of Applied Mathematics, Brown University



# OUTLINE

1. Problem of interest
2. Metastability and accelerated Monte Carlo
3. Performance measure
4. Large deviation properties of empirical measure of metastable diffusion
5. Optimality in two-well model and multi-well model

## **Problem of interest**

---

## RARE EVENT PROBABILITIES AND MCMC

Compute probability  $\mu^\varepsilon(A)$  with respect to a Gibbs measure of the form

$$\mu^\varepsilon(dx) = e^{-V(x)/\varepsilon} dx / Z(\varepsilon),$$

where  $V : \mathbb{R}^d \rightarrow \mathbb{R}$  is the potential of a complex physical system,  $\varepsilon$  is the temperature of the system, and  $A$  does not contain the global minimum of  $V$ .

## RARE EVENT PROBABILITIES AND MCMC

Compute probability  $\mu^\varepsilon(A)$  with respect to a Gibbs measure of the form

$$\mu^\varepsilon(dx) = e^{-V(x)/\varepsilon} dx / Z(\varepsilon),$$

where  $V : \mathbb{R}^d \rightarrow \mathbb{R}$  is the potential of a complex physical system,  $\varepsilon$  is the temperature of the system, and  $A$  does not contain the global minimum of  $V$ .

**Well-known:**  $\mu^\varepsilon(dx)$  is the unique invariant distribution of the diffusion process  $\{X(t)\}_t$  satisfying

$$dX(t) = -\nabla V(X(t)) dt + \sqrt{2\varepsilon} dW(t).$$

# RARE EVENT PROBABILITIES AND MCMC

Compute probability  $\mu^\varepsilon(A)$  with respect to a Gibbs measure of the form

$$\mu^\varepsilon(dx) = e^{-V(x)/\varepsilon} dx / Z(\varepsilon),$$

where  $V : \mathbb{R}^d \rightarrow \mathbb{R}$  is the potential of a complex physical system,  $\varepsilon$  is the temperature of the system, and  $A$  does not contain the global minimum of  $V$ .

**Well-known:**  $\mu^\varepsilon(dx)$  is the unique invariant distribution of the diffusion process  $\{X(t)\}_t$  satisfying

$$dX(t) = -\nabla V(X(t)) dt + \sqrt{2\varepsilon} dW(t).$$

## Markov Chain Monte Carlo (MCMC)

The empirical measure over a large time  $T$ :

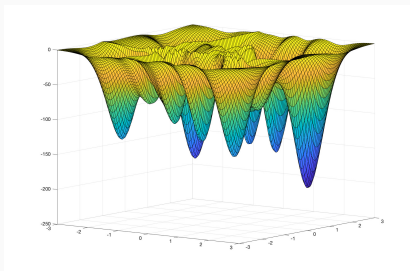
$$\lambda^T(dx) = \frac{1}{T} \int_0^T \delta_{X(t)}(dx) dt \in \mathcal{P}(\mathbb{R}^d).$$

Use  $\lambda^T(A)$  for some large  $T$  as an estimate of  $\mu^\varepsilon(A)$ .

# Metastability

---

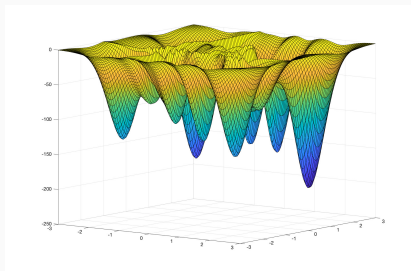
# EXPONENTIAL EXIT TIME



In general  $V$  contains many deep and shallow local minima.



# EXPONENTIAL EXIT TIME



In general  $V$  contains many deep and shallow local minima.

**Exponential exit time:** Mean transition time from one local minimum to another is roughly  $\exp(h/\varepsilon)$  when the temperature  $\varepsilon$  is small, where  $h$  is the barrier height.

## PARALLEL TEMPERING (TWO TEMPERATURES)

Besides  $\varepsilon_1 = \varepsilon$ , introduce higher temperature  $\varepsilon_2 = \varepsilon/\alpha$  with  $\alpha \in (0, 1)$ .

$$dX_1 = -\nabla V(X_1)dt + \sqrt{2\varepsilon_1}dW_1$$

$$dX_2 = -\nabla V(X_2)dt + \sqrt{2\varepsilon_2}dW_2,$$

with  $W_1$  and  $W_2$  independent. Then allow "swaps" with rate

$$ag(x_1, x_2) = a \left( 1 \wedge e^{-\left[ \frac{V(x_1)}{\varepsilon_1} + \frac{V(x_2)}{\varepsilon_2} \right] + \left[ \frac{V(x_2)}{\varepsilon_1} + \frac{V(x_1)}{\varepsilon_2} \right]} \right).$$

## PARALLEL TEMPERING (TWO TEMPERATURES)

Besides  $\varepsilon_1 = \varepsilon$ , introduce higher temperature  $\varepsilon_2 = \varepsilon/\alpha$  with  $\alpha \in (0, 1)$ .

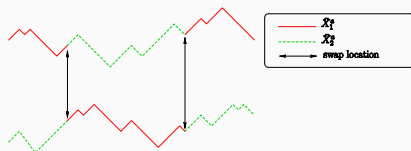
$$dX_1 = -\nabla V(X_1)dt + \sqrt{2\varepsilon_1}dW_1$$

$$dX_2 = -\nabla V(X_2)dt + \sqrt{2\varepsilon_2}dW_2,$$

with  $W_1$  and  $W_2$  independent. Then allow "swaps" with rate

$$ag(x_1, x_2) = a \left( 1 \wedge e^{-\left[\frac{V(x_1)}{\varepsilon_1} + \frac{V(x_2)}{\varepsilon_2}\right] + \left[\frac{V(x_2)}{\varepsilon_1} + \frac{V(x_1)}{\varepsilon_2}\right]} \right).$$

**Particle swapped process:**  $(X_1^a, X_2^a)$



$\mu^{\varepsilon_1}(dx_1)\mu^{\varepsilon_2}(dx_2)$  is the unique invariant distribution of  $(X_1^a, X_2^a)$ .

## INFINITE SWAPPING PROCESS (TWO TEMPERATURES)

**INS process (limit process as swap rate  $a \rightarrow \infty$ ):**

$$dY_1 = -\nabla V(Y_1)dt + \sqrt{2\varepsilon_1\rho(Y_1, Y_2) + 2\varepsilon_2\rho(Y_2, Y_1)}dW_1$$

$$dY_2 = -\nabla V(Y_2)dt + \sqrt{2\varepsilon_2\rho(Y_1, Y_2) + 2\varepsilon_1\rho(Y_2, Y_1)}dW_2,$$

where

$$\rho(x_1, x_2) = e^{-\left[\frac{V(x_1)}{\varepsilon_1} + \frac{V(x_2)}{\varepsilon_2}\right]} / Z_\rho(x_1, x_2),$$

$$Z_\rho(x_1, x_2) = e^{-\left[\frac{V(x_1)}{\varepsilon_1} + \frac{V(x_2)}{\varepsilon_2}\right]} + e^{-\left[\frac{V(x_2)}{\varepsilon_1} + \frac{V(x_1)}{\varepsilon_2}\right]}.$$

The unique invariant distribution of  $(Y_1, Y_2)$  becomes

$$[\mu^{\varepsilon_1}(dx_1)\mu^{\varepsilon_2}(dx_2) + \mu^{\varepsilon_2}(dx_1)\mu^{\varepsilon_1}(dx_2)]/2.$$

**Weighted empirical measure:**

$$\eta^T(dx) = \frac{1}{T} \int_0^T [\rho(Y_1, Y_2)\delta_{(Y_1, Y_2)} + \rho(Y_2, Y_1)\delta_{(Y_2, Y_1)}] dt,$$

Use  $\eta^T(A \times \mathbb{R}^d)$  as an estimate of  $\mu^\varepsilon(A)$ .

# K-TEMPERATURE INS ALGORITHM

$K$ -temperature INS process  $\{X^\varepsilon(t)\}_{t \geq 0} = \{(X_1^\varepsilon(t), \dots, X_K^\varepsilon(t))\}_{t \geq 0}$  for a given  $\alpha = (\alpha_1, \alpha_2, \dots, \alpha_K)$  with  $1 = \alpha_1 \geq \alpha_2 \geq \dots \geq \alpha_K$ :

$$\begin{cases} dX_1^\varepsilon = -\nabla V(X_1^\varepsilon) dt + \sqrt{2\varepsilon} \sqrt{\rho_{11}^\varepsilon/\alpha_1 + \rho_{12}^\varepsilon/\alpha_2 + \dots + \rho_{1K}^\varepsilon/\alpha_K} dW_1 \\ \vdots \\ dX_K^\varepsilon = -\nabla V(X_K^\varepsilon) dt + \sqrt{2\varepsilon} \sqrt{\rho_{K1}^\varepsilon/\alpha_1 + \rho_{K2}^\varepsilon/\alpha_2 + \dots + \rho_{KK}^\varepsilon/\alpha_K} dW_K \end{cases},$$

where

$$\rho_{ij}^\varepsilon \doteq \sum_{\sigma: \sigma(j)=i} w^\varepsilon(x_\sigma; \alpha), \quad w^\varepsilon(x; \alpha) \doteq \frac{\exp[-\frac{1}{\varepsilon} \sum_{\ell=1}^K \alpha_\ell V(x_\ell)]}{\sum_{\sigma \in \Sigma_K} \exp[-\frac{1}{\varepsilon} \sum_{\ell=1}^K \alpha_\ell V(x_{\sigma(\ell)})]}.$$

INS estimator of  $\mu^\varepsilon(A)$  is defined as

$$\theta_{\text{INS}}^{\varepsilon, T} \doteq \frac{1}{T} \int_0^T \sum_{\sigma \in \Sigma_K} w^\varepsilon(X_\sigma^\varepsilon(t); \alpha) 1_A(X_{\sigma(1)}^\varepsilon(t)) dt.$$

# K-TEMPERATURE INS ALGORITHM

$K$ -temperature INS process  $\{X^\varepsilon(t)\}_{t \geq 0} = \{(X_1^\varepsilon(t), \dots, X_K^\varepsilon(t))\}_{t \geq 0}$  for a given  $\alpha = (\alpha_1, \alpha_2, \dots, \alpha_K)$  with  $1 = \alpha_1 \geq \alpha_2 \geq \dots \geq \alpha_K$ :

$$\begin{cases} dX_1^\varepsilon = -\nabla V(X_1^\varepsilon) dt + \sqrt{2\varepsilon} \sqrt{\rho_{11}^\varepsilon/\alpha_1 + \rho_{12}^\varepsilon/\alpha_2 + \dots + \rho_{1K}^\varepsilon/\alpha_K} dW_1 \\ \vdots \\ dX_K^\varepsilon = -\nabla V(X_K^\varepsilon) dt + \sqrt{2\varepsilon} \sqrt{\rho_{K1}^\varepsilon/\alpha_1 + \rho_{K2}^\varepsilon/\alpha_2 + \dots + \rho_{KK}^\varepsilon/\alpha_K} dW_K \end{cases},$$

where

$$\rho_{ij}^\varepsilon \doteq \sum_{\sigma: \sigma(j)=i} w^\varepsilon(x_\sigma; \alpha), \quad w^\varepsilon(x; \alpha) \doteq \frac{\exp[-\frac{1}{\varepsilon} \sum_{\ell=1}^K \alpha_\ell V(x_\ell)]}{\sum_{\sigma \in \Sigma_K} \exp[-\frac{1}{\varepsilon} \sum_{\ell=1}^K \alpha_\ell V(x_{\sigma(\ell)})]}.$$

INS estimator of  $\mu^\varepsilon(A)$  is defined as

$$\theta_{\text{INS}}^{\varepsilon, T} \doteq \frac{1}{T} \int_0^T \sum_{\sigma \in \Sigma_K} w^\varepsilon(X_\sigma^\varepsilon(t); \alpha) 1_A(X_{\sigma(1)}^\varepsilon(t)) dt.$$

**Question:** How to choose  $\alpha$ ?

# Performance measure

---

## UNBIASEDNESS AND DECAY RATE OF VARIANCE

**Time scale:** Good estimation requires  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > 0$ .

### DEFINITION

An estimator  $\theta^{\varepsilon, T^\varepsilon}$  of  $\mu^\varepsilon(A)$  is called **essentially unbiased** if there is  $c_0 \in (0, \infty)$  such that

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left| E\theta^{\varepsilon, T^\varepsilon} - \mu^\varepsilon(A) \right| \geq \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A) + c_0.$$

### DEFINITION

The **decay rate of the variance (per unit time)** of  $\theta^{\varepsilon, T^\varepsilon}$  is defined as

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right).$$



## UNBIASEDNESS AND DECAY RATE OF VARIANCE

**Time scale:** Good estimation requires  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > 0$ .

### DEFINITION

An estimator  $\theta^{\varepsilon, T^\varepsilon}$  of  $\mu^\varepsilon(A)$  is called **essentially unbiased** if there is  $c_0 \in (0, \infty)$  such that

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left| E\theta^{\varepsilon, T^\varepsilon} - \mu^\varepsilon(A) \right| \geq \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A) + c_0.$$

### DEFINITION

The **decay rate of the variance (per unit time)** of  $\theta^{\varepsilon, T^\varepsilon}$  is defined as

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right).$$

- ◇ Performance benchmark is  $2 \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A)$ .

## UNBIASEDNESS AND DECAY RATE OF VARIANCE

**Time scale:** Good estimation requires  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > 0$ .

### DEFINITION

An estimator  $\theta^{\varepsilon, T^\varepsilon}$  of  $\mu^\varepsilon(A)$  is called **essentially unbiased** if there is  $c_0 \in (0, \infty)$  such that

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left| E\theta^{\varepsilon, T^\varepsilon} - \mu^\varepsilon(A) \right| \geq \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A) + c_0.$$

### DEFINITION

The **decay rate of the variance (per unit time)** of  $\theta^{\varepsilon, T^\varepsilon}$  is defined as

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right).$$

- ◇ Performance benchmark is  $2 \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A)$ .
- ◇ Not the best possible decay rate, but the best practically achievable decay rate.

## UNBIASEDNESS AND DECAY RATE OF VARIANCE

**Time scale:** Good estimation requires  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > 0$ .

### DEFINITION

An estimator  $\theta^{\varepsilon, T^\varepsilon}$  of  $\mu^\varepsilon(A)$  is called **essentially unbiased** if there is  $c_0 \in (0, \infty)$  such that

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left| E\theta^{\varepsilon, T^\varepsilon} - \mu^\varepsilon(A) \right| \geq \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A) + c_0.$$

### DEFINITION

The **decay rate of the variance (per unit time)** of  $\theta^{\varepsilon, T^\varepsilon}$  is defined as

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right).$$

- ◇ Performance benchmark is  $2 \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A)$ .
- ◇ Not the best possible decay rate, but the best practically achievable decay rate.
- ◇ Optimize decay rate among essentially unbiased estimators.

## UNBIASEDNESS AND DECAY RATE OF VARIANCE

**Time scale:** Good estimation requires  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > 0$ .

### DEFINITION

An estimator  $\theta^{\varepsilon, T^\varepsilon}$  of  $\mu^\varepsilon(A)$  is called **essentially unbiased** if there is  $c_0 \in (0, \infty)$  such that

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left| E\theta^{\varepsilon, T^\varepsilon} - \mu^\varepsilon(A) \right| \geq \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A) + c_0.$$

### DEFINITION

The **decay rate of the variance (per unit time)** of  $\theta^{\varepsilon, T^\varepsilon}$  is defined as

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right).$$

- ◇ Performance benchmark is  $2 \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(A)$ .
- ◇ Not the best possible decay rate, but the best practically achievable decay rate.
- ◇ Optimize decay rate among essentially unbiased estimators.
- ◇ Conflict between improving the decay rate and achieving essential unbiasedness is insignificant.

# LD properties

---

## SMALL-NOISE DIFFUSION AND QUASIPOTENTIAL

Consider  $\{X_t^\varepsilon\}_{0 \leq t \leq T}$  satisfies

$$dX_t^\varepsilon = b(X_t^\varepsilon)dt + \sqrt{\varepsilon}\sigma(X_t^\varepsilon)dW_t, \quad X_0^\varepsilon = x.$$

Let  $\{O_i\}_{i \in L}$  be all the equilibrium points of  $\dot{x}_t = b(x_t)$  and  $\{X_t^\varepsilon\}$  has an unique invariant distribution  $\mu^\varepsilon$  satisfying

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(O_1) < \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(O_i).$$

Under some conditions,  $\{X_t^\varepsilon\}$  satisfies a large deviation principle with rate function  $I_T$  for any  $T \in (0, \infty)$ .

## SMALL-NOISE DIFFUSION AND QUASIPOTENTIAL

Consider  $\{X_t^\varepsilon\}_{0 \leq t \leq T}$  satisfies

$$dX_t^\varepsilon = b(X_t^\varepsilon)dt + \sqrt{\varepsilon}\sigma(X_t^\varepsilon)dW_t, \quad X_0^\varepsilon = x.$$

Let  $\{O_i\}_{i \in L}$  be all the equilibrium points of  $\dot{x}_t = b(x_t)$  and  $\{X_t^\varepsilon\}$  has an unique invariant distribution  $\mu^\varepsilon$  satisfying

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(O_1) < \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(O_i).$$

Under some conditions,  $\{X_t^\varepsilon\}$  satisfies a large deviation principle with rate function  $I_T$  for any  $T \in (0, \infty)$ . The quasipotential is as

$$Q(x, y) \doteq \inf \{I_T(\phi) : \phi(0) = x, \phi(T) = y, T < \infty\}.$$

# SMALL-NOISE DIFFUSION AND QUASIPOTENTIAL

Consider  $\{X_t^\varepsilon\}_{0 \leq t \leq T}$  satisfies

$$dX_t^\varepsilon = b(X_t^\varepsilon)dt + \sqrt{\varepsilon}\sigma(X_t^\varepsilon)dW_t, \quad X_0^\varepsilon = x.$$

Let  $\{O_i\}_{i \in L}$  be all the equilibrium points of  $\dot{x}_t = b(x_t)$  and  $\{X_t^\varepsilon\}$  has an unique invariant distribution  $\mu^\varepsilon$  satisfying

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(O_1) < \lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(O_i).$$

Under some conditions,  $\{X_t^\varepsilon\}$  satisfies a large deviation principle with rate function  $I_T$  for any  $T \in (0, \infty)$ . The quasipotential is as

$$Q(x, y) \doteq \inf \{I_T(\phi) : \phi(0) = x, \phi(T) = y, T < \infty\}.$$

## DEFINITION

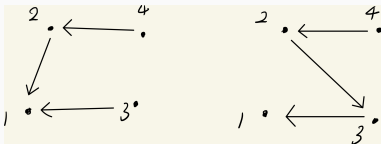
Given a subset  $W \subset L$ , a directed graph consisting of arrows  $i \rightarrow j$  ( $i \in L \setminus W, j \in L, i \neq j$ ) is called a  **$W$ -graph on  $L$**  if

1. every point  $i \in L \setminus W$  is the initial point of exactly one arrow.
2. for any point  $i \in L \setminus W$ , there exists a sequence of arrows leading from  $i$  to some point in  $W$ .



# W-GRAPHS

**Example:**  $L = \{1, 2, 3, 4\}$  and  $W = \{1\}$ .



Denote the set of all  $W$ -graphs by  $G(W)$ .

## DEFINITION

For all  $i \in L$ ,

$$W(O_i) \doteq \min_{g \in G(i)} \left[ \sum_{(m \rightarrow n) \in g} Q(O_m, O_n) \right]$$

and

$$W(O_1 \cup O_i) \doteq \min_{g \in G(1,i)} \left[ \sum_{(m \rightarrow n) \in g} Q(O_m, O_n) \right].$$

# GENERALIZATION OF FREIDLIN-WENTZELL

Freidlin-Wentzell proved that

$$\lim_{\varepsilon \rightarrow 0} -\varepsilon \log \mu^\varepsilon(B_\delta(x)) = W(x) - W(O_1),$$

where  $W(x) \doteq \min_{i \in L} [W(O_i) + Q(O_i, x)]$ .

## THEOREM (DUPUIS AND WU, 2020)

Let  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > h \vee w$ . Given a continuous function  $f : \mathbb{R}^d \rightarrow \mathbb{R}$  and any compact set  $A \subset \mathbb{R}^d$ ,

$$\begin{aligned} \liminf_{\varepsilon \rightarrow 0} -\varepsilon \log & \left| E \left( \frac{1}{T^\varepsilon} \int_0^{T^\varepsilon} e^{-\frac{1}{\varepsilon}f(X_t^\varepsilon)} \mathbf{1}_A(X_t^\varepsilon) dt \right) - \int_{\mathbb{R}^d} e^{-\frac{1}{\varepsilon}f(x)} \mathbf{1}_A(x) \mu^\varepsilon(dx) \right| \\ & \geq \inf_{x \in A} [f(x) + W(x)] - W(O_1) + c - (h \vee w), \end{aligned}$$

with  $h \doteq \min_{i \in L \setminus \{1\}} Q(O_1, O_i)$  and  $w \doteq W(O_1) - \min_{i \in L \setminus \{1\}} W(O_1 \cup O_i)$ .

# DECAY RATE OF VARIANCE

## THEOREM (DUPUIS AND WU, 2020)

Under the same conditions,

$$\begin{aligned} \liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left( T^\varepsilon \cdot \text{Var} \left( \frac{1}{T^\varepsilon} \int_0^{T^\varepsilon} e^{-\frac{1}{\varepsilon} f(X_t^\varepsilon)} 1_A(X_t^\varepsilon) dt \right) \right) \\ \geq \min_{i \in L} \left( R_i^{(1)} \wedge R_i^{(2)} \wedge R_i^{(3)} \right), \end{aligned}$$

where  $R_i^{(1)} \doteq \inf_{x \in A} [2f(x) + Q(O_i, x)] + W(O_i) - W(O_1)$ ,

$$R_1^{(2)} \doteq 2 \inf_{x \in A} [f(x) + Q(O_1, x)] - h,$$

and for  $i \in L \setminus \{1\}$

$$R_i^{(2)} \doteq 2 \inf_{x \in A} [f(x) + Q(O_i, x)] + W(O_i) - 2W(O_1) + W(O_1 \cup O_i),$$

$$R_i^{(3)} \doteq 2 \inf_{x \in A} [f(x) + Q(O_i, x)] + 2W(O_i) - 2W(O_1) - w.$$

# Optimality

---

# DOUBLE WELL

## THEOREM (DUPUIS AND WU, 2020)

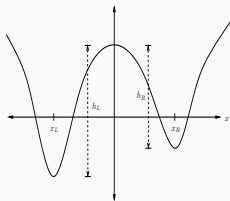
$\theta_{\text{INS}}^{\varepsilon, T^\varepsilon}$  is an essentially unbiased estimator of  $\mu^\varepsilon(A)$ . Moreover,

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta_{\text{INS}}^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right) \geq \begin{cases} r_1(\alpha) \wedge r_3(\alpha), & \text{if } A \subset (-\infty, 0] \\ r_1(\alpha) \wedge r_2(\alpha), & \text{if } A \subset [0, \infty) \end{cases},$$

where  $r_3(\alpha) \doteq 2V(A) - \alpha_K h_L$  with  $V(A) \doteq \inf_{x \in A} V(x)$  and

$$r_1(\alpha) \doteq \inf_{x \in A \times \mathbb{R}^{K-1}} \left[ 2 \sum_{\ell=1}^K \alpha_\ell V(x_\ell) - \min_{\sigma \in \Sigma_K} \left\{ \sum_{\ell=1}^K \alpha_\ell V(x_{\sigma(\ell)}) \right\} \right],$$

$$r_2(\alpha) \doteq \min_{i \in \{2, \dots, K+1\}} \left\{ 2V(A) + \left[ \sum_{\ell=1}^{i-2} \alpha_{K-\ell+1} - \alpha_{K-i+2} \right] (h_L - h_R) \right\} - \alpha_K h_R.$$



# DOUBLE WELL

## THEOREM (DUPUIS AND WU, 2020)

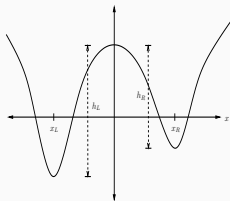
$\theta_{\text{INS}}^{\varepsilon, T^\varepsilon}$  is an essentially unbiased estimator of  $\mu^\varepsilon(A)$ . Moreover,

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta_{\text{INS}}^{\varepsilon, T^\varepsilon} \right) T^\varepsilon \right) \geq \begin{cases} r_1(\alpha) \wedge r_3(\alpha), & \text{if } A \subset (-\infty, 0] \\ r_1(\alpha) \wedge r_2(\alpha), & \text{if } A \subset [0, \infty) \end{cases},$$

where  $r_3(\alpha) \doteq 2V(A) - \alpha_K h_L$  with  $V(A) \doteq \inf_{x \in A} V(x)$  and

$$r_1(\alpha) \doteq \inf_{x \in A \times \mathbb{R}^{K-1}} \left[ 2 \sum_{\ell=1}^K \alpha_\ell V(x_\ell) - \min_{\sigma \in \Sigma_K} \left\{ \sum_{\ell=1}^K \alpha_\ell V(x_{\sigma(\ell)}) \right\} \right],$$

$$r_2(\alpha) \doteq \min_{i \in \{2, \dots, K+1\}} \left\{ 2V(A) + \left[ \sum_{\ell=1}^{i-2} \alpha_{K-\ell+1} - \alpha_{K-i+2} \right] (h_L - h_R) \right\} - \alpha_K h_R.$$



- Optimal  $\alpha^* = (1, 1/2, \dots, (1/2)^{K-2}, \alpha_K^*)$ , where  $\alpha_K^*$  is determined by  $V(A), h_L$  and  $h_R$ .

# DOUBLE WELL

## THEOREM (DUPUIS AND WU, 2020)

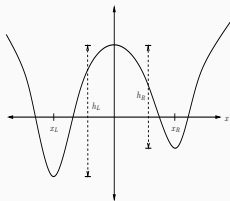
$\theta_{\text{INS}}^{\varepsilon, T^{\varepsilon}}$  is an essentially unbiased estimator of  $\mu^{\varepsilon}(A)$ . Moreover,

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var} \left( \theta_{\text{INS}}^{\varepsilon, T^{\varepsilon}} \right) T^{\varepsilon} \right) \geq \begin{cases} r_1(\alpha) \wedge r_3(\alpha), & \text{if } A \subset (-\infty, 0] \\ r_1(\alpha) \wedge r_2(\alpha), & \text{if } A \subset [0, \infty) \end{cases},$$

where  $r_3(\alpha) \doteq 2V(A) - \alpha_K h_L$  with  $V(A) \doteq \inf_{x \in A} V(x)$  and

$$r_1(\alpha) \doteq \inf_{x \in A \times \mathbb{R}^{K-1}} \left[ 2 \sum_{\ell=1}^K \alpha_{\ell} V(x_{\ell}) - \min_{\sigma \in \Sigma_K} \left\{ \sum_{\ell=1}^K \alpha_{\ell} V(x_{\sigma(\ell)}) \right\} \right],$$

$$r_2(\alpha) \doteq \min_{i \in \{2, \dots, K+1\}} \left\{ 2V(A) + \left[ \sum_{\ell=1}^{i-2} \alpha_{K-\ell+1} - \alpha_{K-i+2} \right] (h_L - h_R) \right\} - \alpha_K h_R.$$



- Optimal  $\alpha^* = (1, 1/2, \dots, (1/2)^{K-2}, \alpha_K^*)$ , where  $\alpha_K^*$  is determined by  $V(A), h_L$  and  $h_R$ .
- Supremum always  $\geq 2V(A) - (1/2)^{K-2}V(A)$ .

## MULTI-WELL

**THEOREM (DUPUIS AND WU, 2021)**

There exists  $B \in (0, \infty)$  such that the following hold. Consider any  $\alpha$  and let  $T^\varepsilon = e^{\frac{1}{\varepsilon}c}$  for some  $c > \alpha_K B$ . Then  $\theta_{\text{INS}}^{\varepsilon, T^\varepsilon}$  is essentially unbiased, and

$$\liminf_{\varepsilon \rightarrow 0} -\varepsilon \log \left( \text{Var}(\theta_{\text{INS}}^{\varepsilon, T^\varepsilon}) T^\varepsilon \right) \geq r(\alpha) - \alpha_K B,$$

where

$$r(\alpha) \doteq \inf_{x \in A \times \mathbb{R}^{d(K-1)}} \left\{ 2 \sum_{\ell=1}^K \alpha_\ell V(x_\ell) - \min_{\sigma \in \Sigma_K} \left\{ \sum_{\ell=1}^K \alpha_\ell V(x_{\sigma(\ell)}) \right\} \right\}.$$

**THEOREM (DUPUIS AND WU, 2021)**

For any closed set  $A$ , and any  $\alpha_K \in (0, (1/2)^{K-1}]$ ,

$$\sup_{(\alpha_2, \dots, \alpha_{K-1}) \in [\alpha_K, 1]^{K-2}} r(\alpha_1, \alpha_2, \dots, \alpha_{K-1}, \alpha_K) = (2 + \alpha_K - (1/2)^{K-2})V(A).$$

The supremum is achieved at  $(\alpha_1^*, \dots, \alpha_{K-1}^*)$  with  $\alpha_\ell^* = (1/2)^{\ell-1}$  for all  $\ell$ .



## SUMMARY

- “Metastability” present a particular challenge for the design of efficient Monte Carlo methods.
- As such, it is natural to use various asymptotic theories to understand issues of algorithm design.
- Have presented one use of large deviation ideas in the context of infinite swapping (and parallel tempering) algorithms to understand the mechanisms that produce variance reduction.
- INS process with a geometric sequence of temperatures explore landscape in a organized and meaningful way. (Ongoing work)

## REFERENCES

Infinite swapping as a limit of parallel tempering:

- ◇ “On the infinite swapping limit for parallel tempering”, Dupuis, Liu, Plattner and Doll, *SIAM J. on MMS*, 10, 986–1022, 2012.

Large deviation estimates:

- ◇ “Large Deviation Properties of the Empirical Measure of a Metastable Small Noise Diffusion”, Dupuis and **Wu**, *J Theo. Prob.*, 2020.

Analysis of INS algorithm:

- ◇ “Analysis and optimization of certain parallel Monte Carlo methods in the low temperature limit”, Dupuis and **Wu**, submitted, 2021.

[CONTACT INFORMATION](mailto:gjwu@kth.se): Guo-Jhen Wu; gjwu@kth.se